# Parameters estimation of the Viscous Blast-Wave model using Machine learning techniques

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## Why This Approach?

We want to determine a set of model parameters that characterize the system formed in the ultra-relativistic heavy ions by comparing model outputs with the experimental results.

But, It's Not So Simple

- From collision to particle dedication different stages are described by different physics/models.
- Need huge CPU time, sometimes beyond our affordable range.
  - Calculation the centrality dependence of bulk observables requires O(10<sup>4</sup>) minimum-bias events @ O(10<sup>-1</sup>)/ hours.
  - O(10<sup>3</sup>) hours per parameter sample.
  - Statistically significant sample size  $> O(10^6)$ .
  - Total computation time > $O(10^9)$  hours  $\sim O(10^5)$  years.
- Experimental measurements always contain some noise.
- A model often has multiple inputs-outputs and usually carries complex interrelationships.

## We Want To Achieve

→ Minimisation of the CPU time.

→ Taking care of the correlations between parameters.

→ Taking care of the different uncertainties in an efficient way.

→ Finally, want to build a data-driven technic that is model-independent.

## Viscous Blast-Wave Model

A Generalized blast-wave model that includes viscous effects by employing a viscosity-based survival scale for geometrical anisotropies formed in the early stages of relativistic heavy-ion collisions was developed by Amaresh Jaiswal and Volker Koch.

$$\frac{d^2 N}{d^2 p_T \, dy} = \frac{1}{(2\pi)^3} \int_0^R r \, dr \int_0^{2\pi} d\varphi \int_{-\infty}^{\infty} \tau \, d\eta_s \, m_T \cosh(y - \eta_s) f.$$

$$v_n(p_T) \equiv \frac{\int_{-\pi}^{\pi} d\varphi_p \, \cos[n(\varphi_p - \Psi_n)] \, \frac{dN}{dy \, p_T \, dp_T \, d\varphi_p}}{\int_{-\pi}^{\pi} d\varphi_p \, \frac{dN}{dy \, p_T \, dp_T \, d\varphi_p}},$$

### The model has six parameters

- Freeze-out temperature, T<sub>r</sub>
- Radial flow velocity,
- $\eta/s$
- Three other constants : m,  $lpha_0$  and  $\kappa$

# **Machine Learning Overview**





**Construction of Design Points / Latin Hypercube Sampling (LHS)** 

Let, the model is to be evaluated on a set of m training points  $X(x_1,...,X_m)$ , in n-dimensional parameter space.

Now, how to generate efficient design points? , i.e., simultaneously optimize emulator accuracy and computation time.

Possible Strategies.

*i*) Factorial Design: k points in each of n dimensions; k<sup>n</sup> total points.

*For k = 20, n = 5 No. Design Points = 3200000; <i>Fails in high dimensions.* 

*ii*) Random Points : A Monte Carlo approach, is good when the design points are large.

With low design points, often leave large regions with no points.

iii) Latin Hypercube Sampling: The most commonly used sampling algorithm.

A semi-random, space-filling design algorithm.

Ref: Latin Hypercube Sampling By Jaroslav

### Latin Hypercube Sampling (LHS)

# **Two Dimension**

In a statical sampling, a square grid containing sample position is a Latin Square if and only if there is only one sample in each row and column.



## Monte Carlo Vs Latin Hypercube

**MCS:** New sample points are generated without taking into account previously generated sample point.

**LHS**: Remember which rows and columns have already been occupied by a sample point.

We use 250 Latin hypercube design point across the n = 6 dimensional parameters space.

### **An Important Property Of MVN**



### An important property of MVN

If a vector  $X(x_1,...,X_n)$  has a multivariate normal distribution then any subset of X also has a multivariate normal distribution.

$$\mu_{2|1} = \mu_2 + \sigma_{21}\sigma_{11}^{-1}(X_1 - \mu_1)$$
  
$$\sigma_{2|1} = \sigma_{22} - \sigma_{21}\sigma_{11}^{-1}\sigma_{12}$$

#### Machine Learning A Probabilistic Perspective by Kevin P. Murphy

### **Multivarient Normal Distribution**

P-dimensional Gaussian PDF for the random vector  $\mathbf{X} = [X_1 \dots X_p]$  has the following form:

$$N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \sim \frac{1}{(2\pi)^{p/2} |\boldsymbol{\Sigma}|^{1/2}} e^{-(\mathbf{X} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{X} - \boldsymbol{\mu})/2}$$

### 

# Sampling from MVN

We generate a sample from univariate gaussian distribution, N(0,1) by mapping the distribution of its <u>Cumulative</u> form, i.e,  $X_i \sim N(0,1)$ 

By extending this idea, for any  $\,\mu\,$  and  $\,\sigma\,$ 

$$X_i \sim N(\mu, \sigma^2) \sim \mu + \sigma N(0, 1)$$

By extending the idea for the multivariant process we can sample a vector **X** from a multivariant gaussian distribution :  $N(X; \bar{\mu}, \Sigma_{n \times n})$ 

$$X_{i} \begin{bmatrix} x_{1i} \\ \cdot \\ \cdot \\ x_{ni} \end{bmatrix} = \begin{bmatrix} \mu_{1} \\ \cdot \\ \mu_{n} \end{bmatrix} + LN \left( \begin{bmatrix} 0_{1} \\ \cdot \\ \cdot \\ 0_{n} \end{bmatrix}, I_{n \times n} \right) \quad \begin{array}{l} \text{L is the } \underline{Cholesky \ Decomposition} \\ \text{of } \Sigma_{n \times n}, i.e., \Sigma_{n \times n} = LL^{T} \end{array}$$

Ref:https://web.mit.edu/urban or book/www/book/chapter7/7.1.3.html

#### Gaussian Process Regression





## Principal Component Analayis (PCA)

Gaussian processes are fundamentally scalar functions, i.e., it take [X]<sub>mxn</sub> input and gives m output. M Design Points N Input Parameters

# But a model may take $[X]_{mxn}$ input and computes $[Y]_{mxp}$ output. P are the different observables.

An obvious workaround is to use independent GP for each of the p outputs. But, we will dissipate the correlated information between different observables.

### Way out of this is PCA!!

Principal Component Analysis (PCA): A dimensionality reduction technic (without loss of too much information), achieved by redistributing the data (<u>Linear</u> <u>Transformation</u>) in a set of new orthonormal basis.

New basis are the eigenvector of the covariance matrix of output  $Y^TY$ 

Ref: M. E. Tipping and C. M. Bishop, "Mixtures of Probabilistic Principal Component Analyzers

### **Numerical Steps For PCA**

Step 1: Standeration of the model output Matrix  $[Y]_{mxp}$ , i.e. centering and scaling each of the collum of Y to zero mean and unit variance.

Step 2: Eigen Value decomposition of covariance matrix  $\mathbf{Y}^{\mathsf{T}}\mathbf{Y}$ Using <u>Singular Value Decomposition (SVD</u>) of the data matrix  $\mathbf{Y}$ .  $Y = U\Sigma V^{T}$ 

**U and V** are the orthogonal matrices ( $U^T U = I \& V^T V = I$ ) known as left and right <u>Singular Vectors</u> respectively and  $\Sigma$  is a diagonal matrix containing the <u>Singular Values</u>.

$$Y^T Y = (U\Sigma V^T)^T U\Sigma V^T = V\Sigma^2 V^T$$

V is the eigenvectors of  $Y^TY$  and  $\Sigma^2$  contains eigenvalues on the diagonal.

Step 3: PCA transformation of the data, i.e., Z= YV, Z is mxp matrix.

Step 4: Use p independent GP emulators for each of the columns of Z.

Step 5: Transformed back to physical space by Y= VZ.

# **Dimentional Reduction**

#### **Output Matrxi:**

### **Eigenvector Matrix:**

$$Y_{p \times m} = \left[ [y_1]_{p \times 1} \dots [y_m]_{p \times 1} \right]_{1 \times m}$$

$$V_{m \times m} = \left[ [e_1]_{m \times 1} \dots [e_m]_{m \times 1} \right]_{1 \times m}$$
**PCA Transformed Output**



$$Z_{p \times m} = YV = \begin{cases} Z_1 = e_{11}X_1 + e_{12}X_2 + \dots + e_{1m}X_m; \\ Z_2 = e_{21}X_1 + e_{22}X_2 + \dots + e_{2m}X_m; \\ \dots \\ Z_p = e_{p1}X_1 + e_{p2}X_2 + \dots + e_{pm}X_p; \end{cases}$$

If the eigenvalue and corresponding eigenvector are in decreasing order.

$$\lambda_1 > \lambda_2 > \dots \lambda_p.$$

First few principal components contained the maximum Variance or Information.

With cause of the minimum loss of information, we can take only the first few principal components and can neglect the rest.

# **PCA Test**



<u>SW=Shapiro–Wilk test , p>0.5 is normal</u>

#### **Single Centrality**

**PCA Results** 

#### **All Centrality**



# **Calibration**

## Extraction of model parameters by comparing experimental data



X : Parameter Vector and D represent all the collected data model + Experiment

$$\begin{array}{ll} \begin{array}{ll} {}_{\textbf{Experiment Result}} & \mathbf{y}_e = \mathbf{y}_e^{\mathrm{true}} + \boldsymbol{\epsilon}_e, & \boldsymbol{\epsilon}_e \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_e) \end{array} \\ \\ \\ \underline{\textbf{Model Result}} & \mathbf{y}_m(\mathbf{x}) = \mathbf{y}_m^{\mathrm{ideal}}(\mathbf{x}) + \boldsymbol{\epsilon}_m, & \boldsymbol{\epsilon}_m \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_m), \end{array} \end{array}$$

 $\mathbf{y}_e = \mathbf{y}_m(\mathbf{x}_{\star}) + \boldsymbol{\epsilon}, \quad \boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \Sigma), \quad \Sigma = \Sigma_e + \Sigma_m,$ 

### **Likelihood**

$$P(\mathcal{D}|\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^m \det \Sigma}} \exp\left\{-\frac{1}{2}[\mathbf{y}_m(\mathbf{x}) - \mathbf{y}_e]^{\mathrm{T}} \Sigma^{-1}[\mathbf{y}_m(\mathbf{x}) - \mathbf{y}_e]\right\}$$

### **Uniform Prior Distribution**

$$P(\mathbf{x}) \propto \begin{cases} 1 & \text{if } \min(x_i) \le x_i \le \max(x_i) \text{ for all } i \\ 0 & \text{else.} \end{cases}$$

It is better to deal with the log Prior

$$\log P(\mathbf{x}|\mathcal{D}) = \log P(\mathcal{D}|\mathbf{x}) + \log P(\mathbf{x}) + \text{const},$$

$$\log P(\mathcal{D}|\mathbf{x}) = -\frac{1}{2}\mathbf{d}^{\mathrm{T}}\Sigma^{-1}\mathbf{d} - \frac{1}{2}\log(\det\Sigma), \quad \mathbf{d} = \mathbf{y}_m(\mathbf{x}) - \mathbf{y}_e,$$

$$\log P(\mathcal{D}|\mathbf{x}) = -\frac{1}{2}\mathbf{d} \cdot \boldsymbol{\alpha} - \sum_{i} \log L_{ii}. \quad \boldsymbol{\alpha} = \Sigma^{-1}\mathbf{d}.$$

$$\Sigma = LL^{\mathrm{T}}, \quad \boldsymbol{\alpha} = \Sigma^{-1} \mathbf{d}.$$
  
 $LL^{\mathrm{T}} \boldsymbol{\alpha} = \mathbf{d}$ 

We Maximize the log probability using MCMC Algorithm We Run MCMC on

$$P(\mathbf{z}_{\exp}|\mathbf{x}_{\star}) \propto \exp\left\{-\frac{1}{2}(\mathbf{z}_{\star} - \mathbf{z}_{\exp})^{\mathsf{T}} \Sigma_{z}^{-1}(\mathbf{z}_{\star} - \mathbf{z}_{\exp})\right\}$$

**Error Estimation** 

$$\Sigma_e = \Sigma_e^{\text{stat}} + \Sigma_e^{\text{sys}}.$$

Statistical uncertainty are uncorrelated

$$\Sigma_e^{\text{stat}} = \text{diag}\Big[(\sigma_1^{\text{stat}})^2, (\sigma_2^{\text{stat}})^2, \dots, (\sigma_m^{\text{stat}})^2\Big]$$

0

Systematic uncertainty usually has a correlation structure

$$\Sigma_{ij} = \operatorname{cov}(y_i, y_j) = \rho_{ij} \sigma_i \sigma_j \qquad \rho_{ij}^{\text{sys}} = \exp\left[-\frac{1}{2} \left(\frac{c_i - c_j}{\ell}\right)^2\right]$$

<u>C is the midpoint of the pT Bin and I=1</u>

$$\Sigma_m = \Sigma_m^{\text{pred}} + \Sigma_m^{\text{stat}} + \Sigma_m^{\text{sys}}.$$
$$\Sigma_m = \Sigma_m^{\text{GP}} + \Sigma_m^{\text{sys}}.$$

 $\Sigma$  = Predictive Covariance in PC space and V is the PCA transformation matrix

$$\Sigma^{\rm GP}_m = V \Sigma^{\rm GP}_{m,z} V^{\rm T}.$$

MCMC Result





 $(1/2\pi p_T) dN/(dp_T dy) [c^2/GeV^2$ 



# **Programming libraries and Sources**

# **Open-Sources Python Libraires**

- NumPy
- SciPy
- scikit-learn
- H5py
- matplotlib
- emcee

# Other Sources: Reading material

- Bayesian analysis of computer code outputs: A tutorial by A. O'Hagan
- Nonparametric Bayesian Methods (Gaussian Processes) Jun Zhu
- https://arxiv.org/abs/1902.11082.
- An Intuitive Tutorial to Gaussian Processes Regression Jie Wang.
- Bayesian estimation of the specific shear and bulk viscosity of quark– gluon plasma/ Nat. Phys. 15, 1113–1117 (2019).